Lecture 10. Modeling Process and Model Diagnostics

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This Lecture

- Modeling process
- Goodness of fit
- Residuals

Modeling Process





- The choice of a model class is often driven by many factors, including data characteristics, expressiveness, interpretability, computational efficiency...
- If predictive performance (expressiveness) is the main concern
 - try deep neural networks for image/text/speech data.
 - try random forests when high-level features are available.
- GLMs can be good in terms of interpretability.



- More data is often better.
- With right features, even simple models can work well.
- Exploratory analysis can suggest useful features and models.



- Fitting is usually formulated as an optimization problem.
- MLE is often used to learn a statistical model.
- If predictive performance is the main concern, optimize the performance measure directly.
- Sophisticated optimization algorithms may be needed.
 - For GLM, Fisher scoring often works well for MLE.



- Check model assumption
 - Check goodness of fit, residual plot et al on training set.
 - A good fit on the training set may mean overfitting.
- Check predictive performance
 - Check cross-validation score, validation set performance.
- Reconsider model class or data if checks are not satisfactory.



 After checks on the model, the model can then be used to make predictions or draw conclusions (such as significance of variables, variable importance).

Goodness of Fit

Deviance

- Null model
 - Includes only the intercept term in the GLM.
 - Variation in y's comes from the random component only.
- Full model (saturated model)
 - Fit an exponential family distribution for each example.
 - The exponential family distribution for (\mathbf{x}_i, y_i) is $f(y | \text{mean} = y_i)$.
 - Variation in y's comes from the systematic component only.
- GLM
 - Summarizes data with a few parameters.
 - The exponential family distribution for (\mathbf{x}_i, y_i) is $f(y \mid \text{mean} = \hat{\mu}_i)$, where $\hat{\mu}_i = g^{-1}(\mathbf{x}_i^{\top}\hat{\beta})$.

• Scaled deviance

$$D^*(\mathbf{y}; \hat{\boldsymbol{\mu}}) = 2\sum_i \ln f(y_i \mid \text{mean} = y_i) - 2\sum_i \ln f(y_i \mid \text{mean} = \hat{\mu}_i)$$

This is twice the difference between log-likelihood of the full model and the maximum log-likelihood achievable for the GLM.

Deviance

$$D(\mathbf{y}; \hat{\boldsymbol{\mu}}) = b(\phi) D^*(\mathbf{y}; \hat{\boldsymbol{\mu}}).$$

Deviance is thus scaled deviance with the nuisance parameter removed.

Example. Gaussian The scaled deviance is

$$D^{*}(\mathbf{y}; \hat{\boldsymbol{\mu}}) = 2\sum_{i} \left(\ln \frac{1}{\sqrt{2\pi\sigma}} - \frac{(y_{i} - y_{i})^{2}}{2\sigma^{2}} \right) - 2\sum_{i} \left(\ln \frac{1}{\sqrt{2\pi\sigma}} - \frac{(y_{i} - \hat{\mu}_{i})^{2}}{2\sigma^{2}} \right) = \sum_{i} \frac{(y_{i} - \hat{\mu}_{i})^{2}}{\sigma^{2}}.$$

The deviance is

$$D(\mathbf{y}; \hat{\boldsymbol{\mu}}) = \sigma^2 D^*(\mathbf{y}; \hat{\boldsymbol{\mu}}) = \sum_i (y_i - \hat{\mu}_i)^2.$$

distribution	deviance
normal	$\sum (y - \hat{\mu})^2$
Poisson	$2\sum(y\lnrac{y}{\hat{\mu}}-(y-\hat{\mu}))$
binomial	$2\sum(y\ln\frac{y}{\hat{\mu}}+(m-y)\ln\frac{m-y}{m-\hat{\mu}})$
Gamma	$2\sum(-\ln\frac{y}{\hat{\mu}}+\frac{y-\hat{\mu}}{\hat{\mu}})$
inverse Gaussian	$\sum (y - \hat{\mu})^2 / (\hat{\mu}^2 y)$

Recall

```
> logLik(fit.ig.inv)
'log Lik.' -25.33805 (df=5)
> logLik(fit.ig.invquad)
'log Lik.' -50.26075 (df=5)
> logLik(fit.ig.log)
'log Lik.' -45.55859 (df=5)
```

Inverse Gaussian regression with inverse link has the best fit (much better than the other two).

```
summary(fit.ig.inv)
>
   Null deviance: 0.24788404
                              on 17 degrees of freedom
Residual deviance: 0.00097459
                              on 14 degrees of freedom
> summary(fit.ig.invquad)
   Null deviance: 0.24788 on 17
                                  degrees of freedom
Residual deviance: 0.01554 on 14
                                  degrees of freedom
> summary(fit.ig.log)
   Null deviance: 0.2478840 on 17 degrees of freedom
Residual deviance: 0.0092164
                             on 14
                                    degrees of freedom
```

- Inverse link has best fit.
- Same conclusion as obtained by looking at the log-likelihoods.
- summary function provides a comparison with the full model and null model.

Generalized Pearson X² statistic

- Recall: $var(Y) = b(\phi)A''(\eta)$ for a natural exponential family.
- $\operatorname{var}(Y)/b(\phi)$ depends only on η , and thus only on μ .
- Often, $var(Y)/b(\phi)$ is called the variance function $V(\mu)$.
- Pearson X² statistic is

$$X^2 = \sum (y - \hat{\mu})^2 / V(\hat{\mu}),$$

where $V(\hat{\mu})$ is the estimated variance function.

• The scaled version is $X^2/b(\phi)$.

distribution	<i>X</i> ²
normal Poisson	$\sum (y - \hat{\mu})^2$ $\sum (y - \hat{\mu})^2 / \hat{\mu}$
binomial	$\sum \frac{(y-\hat{\mu})^2}{\hat{\mu}(1-\hat{\mu})}$
Gamma	$\sum (y - \hat{\mu})^2) / \hat{\mu}^2$
inverse Gaussian	$\sum (y-\hat{\mu})^2/\hat{\mu}^3$

Asymptotic distribution

- If the model is true, then the scaled deviance or the scaled Pearson X^2 statistic asymptotically follows χ^2_{n-p} , where *n* is the number of examples, and *p* is the number of parameters estimated.
- In principle, this can be used to test goodness of fit, but this does not really work well.
- A test on the scaled deviance or the scaled Pearson X² statistic cannot be used to justify that the model is correct.

Residuals

Response residual

• This is the difference between the output and fitted mean

$$r_R = y - \hat{\mu}.$$

• Measures deviation from systematic effect on an absolute scale.

Pearson residuals

• This is the normalized response residual

$$r_P = \frac{y - \hat{\mu}}{\sqrt{V(\hat{\mu})}}$$

• Constant variance and mean zero if model is correct.

distribution	Pearson residual
normal	$y - \hat{\mu}$
Poisson	$(y-\hat{\mu})/\sqrt{\hat{\mu}}$
binomial	$(y-\hat{\mu})/\sqrt{\hat{\mu}(1-\hat{\mu})}$
Gamma	$(y-\hat{\mu})/\hat{\mu}$
inverse Gaussian	$(y-\hat{\mu})/\hat{\mu}^{3/2}$
	$(J P^{*})/P^{*}$

Working residuals

• Recall: in the IRLS interpretation of Fisher scoring, at each iteration we try to fit the *adjusted response* vector

$$\mathbf{z} = \mathbf{G}\mathbf{y} - \mathbf{G}\boldsymbol{\mu} + \mathbf{X}\boldsymbol{\beta},$$

where $\mathbf{G} = \operatorname{diag}(g'(\mu_1), \ldots, g'(\mu_n)).$

• The adjusted response for (\mathbf{x}, y) is

$$z = g'(\mu)(y - \mu) + \mathbf{x}^{\top} \beta.$$

• The working residual is

$$r_W = z - \xi = (y - \hat{\mu})g'(\mu) = (y - \hat{\mu})\frac{\partial \xi}{\partial \mu}|_{\mu = \hat{\mu}},$$

where $\xi = \mathbf{x}^{\top} \beta$.

Deviance residuals

• This is the signed contribution of each example to the deviance

$$r_D = \operatorname{sign}(y - \hat{\mu})\sqrt{d},$$

where $\sum_i d_i = D$.

- Closer to a normal distribution (less skewed) than Pearson residuals.
- Often better for spotting outliers.

distribution	deviance residual
normal	$y-\hat{\mu}$
Poisson	${ m sign}(y-\hat{\mu})\sqrt{2(y\lnrac{y}{\hat{\mu}}-(y-\hat{\mu}))}$
binomial	$\operatorname{sign}(y-\hat{\mu})\sqrt{2(y\ln rac{y}{\hat{\mu}}+(m-y)\ln rac{m-y}{m-\hat{\mu}})}$
Gamma	${ m sign}(y-\hat{\mu})\sqrt{2(-\lnrac{y}{\hat{\mu}}+rac{y-\hat{\mu}}{\hat{\mu}})}$
inverse Gaussian	$(y-\hat{\mu})/\hat{\mu}\sqrt{y}$

Computing residuals in R

```
> resid(fit.ig.inv, 'response')
> resid(fit.ig.inv, 'pearson')
> resid(fit.ig.inv, 'working')
> resid(fit.ig.inv, 'deviance')
```

What You Need to Know

- Modeling process
- Goodness of fit: deviance and Pearson X^2 statistic
- Response, working, Pearson, and deviance residuals